

Neurofilament light chain level is a weak risk factor for the development of MS

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Supplemental data
 at Neurology.org

ABSTRACT

Objective: To determine the prognostic value of selected biomarkers in clinically isolated syndromes (CIS) for conversion to multiple sclerosis (MS) and disability accrual.

Methods: Data were acquired from 2 CIS cohorts. The screening phase evaluated patients developing clinically definite MS (CIS-CDMS) and patients who remained as CIS during a 2-year minimum follow-up (CIS-CIS). We determined levels of neurofascin, semaphorin 3A, fetuin A, glial fibrillary acidic protein, and neurofilament light (NfL) and heavy chains in CSF (estimated mean [95% confidence interval; CI]). We evaluated associations between biomarker levels, conversion, disability, and magnetic resonance parameters. In the replication phase, we determined NfL levels ($n = 155$) using a 900 ng/L cutoff. Primary endpoints in uni- and multivariate analyses were CDMS and 2010 McDonald MS.

Results: The only biomarker showing significant differences in the screening was NfL (CIS-CDMS 1,553.1 [1,208.7–1,897.5] ng/L and CIS-CIS 499.0 [168.8–829.2] ng/L, $p < 0.0001$). The strongest associations were with brain parenchymal fraction change ($r_s = -0.892$) and percentage brain volume change ($r_s = -0.842$) at 5 years. NfL did not correlate with disability. In the replication phase, more NfL-positive patients, according to the cutoff, evolved to MS. Every 100-ng/L increase in NfL predicted CDMS (hazard ratio [HR] = 1.009, 95% CI 1.005–1.014) and McDonald MS (HR = 1.009, 95% CI 1.005–1.013), remaining significant for CDMS in the multivariate analysis (adjusted HR = 1.005, 95% CI 1.000–1.011). This risk was lower than the presence of oligoclonal bands or T2 lesions.

Conclusions: NfL is a weak independent risk factor for MS. Its role as an axonal damage biomarker may be more relevant as suggested by its association with medium-term brain volume changes.

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GLOSSARY

aHR = adjusted hazard ratio; **BPF** = brain parenchymal fraction; **CDMS** = clinically definite multiple sclerosis; **CI** = confidence interval; **CIS** = clinically isolated syndrome; **CV** = coefficient of variation; **DMT** = disease-modifying treatment; **EDSS** = Expanded Disability Status Scale; **FA** = fetuin A; **Gd** = gadolinium; **GFAP** = glial fibrillary acidic protein; **HR** = hazard ratio; **MS** = multiple sclerosis; **NfH** = neurofilament heavy chain; **NfL** = neurofilament light chain; **OCB** = oligoclonal bands; **PBVC** = percentage brain volume change; **sema3A** = semaphorin 3A; **SIENA** = Structural Image Evaluation, using Normalization, of Atrophy.

Identifying patients with clinically isolated syndromes (CIS) who will present a second attack and determining the degree of disability accrual over the medium- to long-term are crucial goals in multiple sclerosis (MS). Although MRI is the most reliable prognostic marker^{1,2} and oligoclonal bands (OCBs) also have an important role,^{3,4} the disease heterogeneity hinders a more individualized prognosis. Therefore, the search for biomarkers that capture the different aspects

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of MS is still necessary. Although many have been identified, few are currently useful in the clinical setting.⁵ Consequently, we aimed to determine the prognostic value of selected biomarkers for conversion to MS and disability accrual, and to assess their associations with MRI inflammatory activity and neurodegenerative parameters.

METHODS Patients. We acquired longitudinal data from CIS cohorts at the Multiple Sclerosis Center in Vall d'Hebron Hospital, Barcelona,⁶ and the Neurology Department of Ramón y Cajal Hospital, Madrid. Inclusion criteria were age younger than 50 years and baseline clinical evaluation within the first 3 months of disease onset. We recorded baseline demographic and clinical characteristics and conducted follow-up visits every 3 to 6 months, assessing for relapses, clinical worsening, and excluding patients in whom we reached alternative diagnoses. A second attack established the diagnosis of clinically definite MS (CDMS).⁷ Definitions of disability accrual were a sustained increase in the Expanded Disability Status Scale (EDSS) of 1.0 point over 1 year or confirmed EDSS score ≥ 3.0 .

Biological samples. We collected venous blood and CSF within 3 months of disease onset for OCB determination in the 2 hospitals. Remnant samples were aliquoted and stored at -80°C until further use.

MRI acquisition and analysis. At Vall d'Hebron Hospital, brain MRI acquisition on 1.5- or 3.0-tesla superconductive magnets included the following sequences, obtained with a contiguous 3- to 5-mm slice thickness: transverse proton density and T2-weighted conventional or fast spin-echo, transverse and sagittal T2 fluid-attenuated inversion recovery, and unenhanced and contrast-enhanced T1-weighted spin-echo. We included routine baseline spinal cord MRIs in November 2007. Baseline scans were done within 5 months of disease onset and follow-up MRIs at 1 year and every 5 years thereafter.⁶ The Ramón y Cajal cohort followed a similar acquisition protocol, performing baseline and follow-up MRIs within 3 months after the CIS and at 1 year, including spinal cord MRI in case of myelitis.

Neuroradiologists scored T2 lesion number and location, and gadolinium (Gd)-enhancing and new T2 lesion number. T2 lesion volume was calculated using the semiautomated Jim medical image display package (Xinapse Systems, Ltd., Colchester, UK). Brain volume parameters were obtained using the Structural Image Evaluation, using Normalization, of Atrophy (SIENA) software, part of FSL (FMRIB Software Library).⁸ Single-time-point analysis to obtain brain parenchymal fraction (BPF) estimates was performed with SIENAX. Two-time-point analysis was performed with SIENA, estimating the percentage brain volume change (PBVC) between 2 input images of the same subject at different time points.

Biological markers. We selected fetuin A (FA),⁹ semaphorin 3A (sema3A),^{10,11} glial fibrillary acidic protein (GFAP),^{12–15} neurofilament heavy (NfH) and light (NfL) chains,^{13–20} and neurofascin^{21,22} after a PubMed search in 2008, which we updated regularly throughout the study duration, including the terms “biological markers” or “biomarkers” and “multiple sclerosis” or “clinically isolated syndromes.”

Except for NfH, we determined biomarker levels using commercial sandwich ELISA kits according to the manufacturers'

recommendations (FA: BioVendor, Brno, Czech Republic; NfL: UmanDiagnostics AB, Umeå, Sweden; GFAP: Abnova, Taipei City, Taiwan; neurofascin: USCN Life Science Inc., Wuhan, China; and sema3A: USCN Life Science Inc.). We optimized the assays using varying sample dilutions measured in duplicates. NfH levels were determined using an electrochemiluminescence-based solid-phase sandwich immunoassay¹⁷ at the University Hospital Basel, Switzerland, selected among different assays for its higher sensitivity.^{18,23,24} All samples were tested blinded to clinical data and outcome measures. Whenever possible according to availability, we tested samples from the same patients using a different, unfrozen aliquot.

Experimental design. Screening phase. We selected 2 opposite groups with a 2-year minimum follow-up from the Vall d'Hebron cohort: CIS-CDMS (n = from 33 to 38), including patients with 3 to 4 Barkhof-Tintore criteria on baseline MRI and presence of OCBs who converted to CDMS; and CIS-CIS (n = from 33 to 39), comprising patients with 0 Barkhof-Tintore criteria and absence of OCBs who remained as CIS (figure e-1A at Neurology.org).

Replication phase. Only NfL levels showed significant differences in the screening and were thus evaluated in a replication phase. We selected consecutive patients with available CSF samples from the Vall d'Hebron and Ramón y Cajal cohorts (n = 155). Patients from Vall d'Hebron were selected from a different time range (2009–2011) than the NfL screening phase because of sample availability. NfL determination was performed in each hospital separately using the above-mentioned assay (figure e-1B).

Statistical analysis. We performed parametric and nonparametric descriptive statistics depending on the normality of the distributions of the continuous variables.

Screening phase. For each biomarker, we calculated sample size for 80% power at the 0.05 level of significance based on previous data. We used generalized linear models to compare CIS groups in terms of conversion to CDMS and assess biomarker levels and disability accrual, with the Bonferroni correction for multiple comparisons when applicable. Data are reported as the estimated mean (95% confidence interval [CI]). Potential covariates were age, sex, CIS topography, time from CIS to lumbar puncture, sample storage time, and disease-modifying treatment (DMT) before CDMS. We calculated Spearman correlations for MRI inflammatory activity parameters at baseline, 1, and 5 years (number of T2, Gd-enhancing, and new T2 lesions; T2 lesion volume), and partial correlations for neurodegenerative parameters at 1 and 5 years of follow-up (BPF change [BPF Δ] adjusted for age and baseline Gd-enhancing lesions, and PBVC). We calculated BPF Δ by subtracting baseline from 1- and 5-year follow-up SIENAX estimates and dividing by the baseline values, then multiplying by 100. We assessed the PBVC estimated by SIENA for changes from baseline to 1 and 5 years. Finally, we assessed Spearman correlations between NfL and NfH levels in a subgroup of 42 patients with determination of both biomarkers.

Replication phase. Primary endpoints were conversion to CDMS and 2010 McDonald MS according to NfL status (positive/negative) based on a 900 ng/L cutoff value, established as the mean ± 3 SD of NfL levels in a control group of patients with noninflammatory neurologic diseases.²⁵ We performed uni- and multivariate Cox proportional hazard regression analyses for NfL levels as a continuous or dichotomic variable. Covariates were OCBs, baseline T2 lesion number, DMT before MS, and hospital where NfL levels were determined. We used generalized linear models to compare NfL levels in terms of disability accrual (EDSS score ≥ 3.0), and Spearman correlations or generalized linear models to determine associations between NfL and MRI inflammatory parameters at baseline and 1 year. We could not assess brain volume

changes because of the few estimated measures at the time of the analysis.

We performed statistical tests on the 0.05 level of significance, using the IBM SPSS Statistics (version 20.0; IBM Corp., Armonk, NY). We also prepared figures using GraphPad Prism 5.02 for Windows (GraphPad Software, La Jolla, CA).

Standard protocol approvals, registrations, and patient consents. This study received approval from the corresponding local ethical committees in both hospitals and all patients signed a written informed consent.

RESULTS Screening phase. Appendix e-1A describes preliminary analyses. Baseline demographic and clinical characteristics were comparable between CIS groups (table 1) except for the higher frequency of optic neuritis in CIS-CIS regarding FA, NfL, and NfH. The proportion of patients on DMT before CDMS varied between 8.8% and 17.7% for each biomarker. Mean follow-up was longer for CIS-CDMS than CIS-CIS tested for GFAP and NfH. The differences in storage time between CIS-CDMS and CIS-CIS samples for NfH did not correlate with this biomarker's levels ($r_s = -0.131, p = 0.256$). NfH levels correlated with age ($r_s = 0.230, p = 0.044$). We found no differences in time from CIS to lumbar puncture. Therefore, for each protein, we adjusted the analyses for age, topography, and/or DMT.

FA, GFAP, and NfH results were negative (figure e-2, A–C). We could not detect neurofascin and sema3A.

As for NfL levels, we found a significant difference between CIS-CDMS (1,553.1 [1,208.7–1,897.5] ng/L) and CIS-CIS (499.0 [168.8–829.2] ng/L) ($p < 0.0001$) (figure 1). The average intra- and interplate coefficients of variation (CVs) were 9.7% and 0.5%, respectively. We found no significant differences for EDSS progression of 1 point or ≥ 3.0 , even though in the latter case, NfL levels were 1,506.1 (645.2–2,367.1) ng/L in the 15 patients with EDSS score ≥ 3.0 and 906.6 (693.2–1,120.1) ng/L if < 3.0 ($p = 0.185$). We found significant correlations between NfL levels and inflammatory activity parameters on MRI, except for Gd-enhancing lesions at 5 years (table 2). As for neurodegenerative parameters, NfL and both BPF Δ and PBVC at 5 years showed the strongest correlations (table 2, figure 2).

Finally, we found a correlation between NfL and NfH levels ($r_s = 0.466, p = 0.002$).

Replication phase. Appendix e-1B describes preliminary analyses. Baseline demographic and clinical characteristics are shown in table e-1. Median (interquartile range) NfL levels were 1,238.3 (1,782.1) ng/L for patients converting to CDMS and 555.8 (825.5) ng/L for CIS-CIS. Average intra- and interplate CV were 4.9% and 7.0% at Vall d'Hebron ($n = 93$) and 5.1% and 8.8% at Ramón y Cajal ($n = 62$).

According to the cutoff, 63 patients (40.6%) were NfL positive. Significantly more NfL-positive patients

Table 1 Demographic and clinical characteristics

| | FA and NfL (2010) ^a | | GFAP (2011) | | NfH (2012) | |
|--|--------------------------------|-------------|---------------|---------------|--------------|--------------|
| | CIS-CDMS | CIS-CIS | CIS-CDMS | CIS-CIS | CIS-CDMS | CIS-CIS |
| No. | 35 | 33 | 33 | 33 | 38 | 39 |
| Mean age (SD), y | 29.3 (7.3) | 30.1 (9.7) | 30.7 (8.0) | 30.4 (8.2) | 30.7 (7.5) | 31.6 (8.2) |
| Females, n (%) | 24 (68.6) | 26 (78.8) | 21 (63.6) | 24 (72.7) | 28 (73.7) | 32 (82.1) |
| Mean follow-up (SD), ^b mo | 78.8 (33.4) | 57.9 (35.2) | 79.1 (36.2) | 43.7 (50.0) | 104.3 (31.1) | 74.3 (36.4) |
| CIS topography, n (%) ^c | | | | | | |
| Optic nerve | 10 (28.6) | 22 (66.7) | 11 (33.3) | 20 (60.6) | 10 (26.3) | 24 (61.5) |
| Brainstem | 9 (25.7) | 5 (15.2) | 7 (21.2) | 7 (21.2) | 8 (21.1) | 4 (10.3) |
| Spinal cord | 11 (31.4) | 4 (12.1) | 10 (30.3) | 4 (12.1) | 14 (36.8) | 3 (7.7) |
| Other | 5 (14.3) | 2 (6.1) | 5 (15.2) | 2 (6.1) | 6 (15.8) | 8 (20.5) |
| Mean time from CIS to lumbar puncture (SD), d | 58.9 (53.5) | 80.1 (72.3) | 184.4 (534.1) | 199.7 (303.6) | 57.1 (52.5) | 81.1 (169.2) |
| Frozen samples mean storage time (SD), ^d mo | 88.5 (34.1) | 80.2 (33.5) | 91.5 (34.5) | 75.0 (44.3) | 113.1 (33.7) | 97.6 (40.5) |

Abbreviations: CDMS = clinically definite multiple sclerosis; CIS = clinically isolated syndrome; FA = fetuin A; GFAP = glial fibrillary acidic protein; NfH = neurofilament heavy chain; NfL = neurofilament light chain.

The screening phase was performed between 2010 and 2012. Categorical variables: χ^2 test. Continuous variables: Student t test for 2 groups, analysis of variance for multiple comparisons.

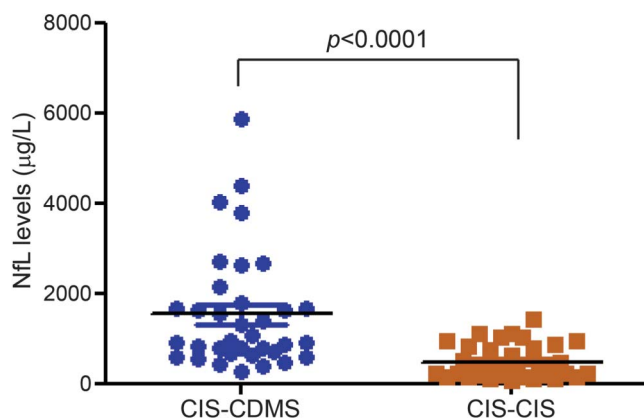
^a Years in parentheses show the time when each biomarker was tested. Samples were obtained between 1995 and 2008 for FA, NfL, and GFAP, and between 1995 and 2009 for NfH. CSF samples were aliquoted and stored at -80°C until their use.

^b FA and NfL: $p = 0.050$; GFAP: $p = 0.002$; NfH: $p < 0.0001$.

^c FA and NfL: $p = 0.017$; NfH: $p = 0.002$; GFAP: $p = 0.099$.

^d NfH levels: $p = 0.037$.

Figure 1 NfL levels in the 2 CIS groups



Results adjusted for CIS topography and disease-modifying treatment. CDMS = clinically definite multiple sclerosis; CIS = clinically isolated syndrome; NfL = neurofilament light chain.

converted to CDMS or fulfilled McDonald MS over time (data not shown).

When evaluating NfL levels as a continuous variable in the univariate analyses, we observed a 1-point increase in risk of both CDMS and McDonald MS for every 100-ng/L increment. The adjusted hazard ratio (aHR) for CDMS remained significant in the multivariate analysis (table 3).

As for NfL status, NfL positivity was a predictor of CDMS (HR 2.279, 95% CI 1.283–4.049, $p = 0.005$) and McDonald MS (HR 2.801, 95% CI 1.761–4.455, $p < 0.0001$) in the univariate analyses; however, this significance was lost in the multivariate models (aHR 1.220, 95% CI 0.656–2.269, $p = 0.530$ and aHR 1.347, 95% CI 0.809–2.244, $p = 0.252$, respectively).

Regarding disability, although NfL levels were 2,635.6 (1,226.4–4,044.8) ng/L in the 11 patients with EDSS score ≥ 3.0 and 1,560.5 (934.8–2,186.2) ng/L if < 3.0 , this difference was not statistically significant ($p = 0.172$).

Finally, we found significant correlations between NfL levels and T2 lesion volume and Gd-enhancing lesion number on baseline MRIs ($n = 63$, $r_s = 0.517$,

$p < 0.0001$ and $n = 146$, $r_s = 0.469$, $p < 0.0001$, respectively) and follow-up MRIs ($n = 36$, $r_s = 0.420$, $p = 0.011$ for new T2 lesion volume and $n = 94$, $r_s = 0.231$, $p = 0.025$ for Gd-enhancing lesions). Higher NfL levels were associated with a higher T2 lesion number at baseline and 1 year (data not shown).

DISCUSSION In the screening phase, we aimed to identify differences in biomarker levels under the rationale that, if none were found between 2 opposite CIS groups,²⁶ none would be found in consecutive patients, a scenario that better resembles clinical practice. In addition, this phase allows a strict use of biological samples. Only NfL levels were significantly higher in patients who converted to MS compared to nonconverters. Neurofilaments are type IV intermediate filaments specific for neurons, released into the CSF when axonal damage occurs.^{16,27,28} A study with 38 patients reported higher NfL levels in CIS-CDMS compared to CIS-CIS,¹⁸ a finding unconfirmed in other investigations including 9 to 36 patients per group.^{20,29,30} Such discrepancies could be attributable to a small sample size in the converter group,^{20,30} homogeneity in baseline inflammatory parameters such as proportion of OCBs present in both CIS groups,²⁹ or the method used to detect NfL.³⁰ In addition, sample size calculations are not referred and some groups may be too small to properly demonstrate any differences. Our study was adequately powered and thus confirms the differences with a well-balanced distribution between groups.

We found no associations between NfL levels and disability by the EDSS. Except for one, most studies showed significant correlations, but the analyses included CIS and other MS phenotypes, observing the strongest correlations in progressive forms.^{13,14,18,31–33} Khalil et al.²⁰ evaluated 67 patients with CIS exclusively and found a marginal correlation coefficient of $r_s = 0.324$ ($p < 0.05$). When assessing NfL levels according to EDSS score ≥ 3.0 in both phases, they were higher in patients with more severe disability but the difference was not significant, a finding probably influenced by the few patients who reached such EDSS score.

Table 2 Correlations between baseline neurofilament light chain levels and MRI measures at baseline, 1 year, and 5 years: Screening phase

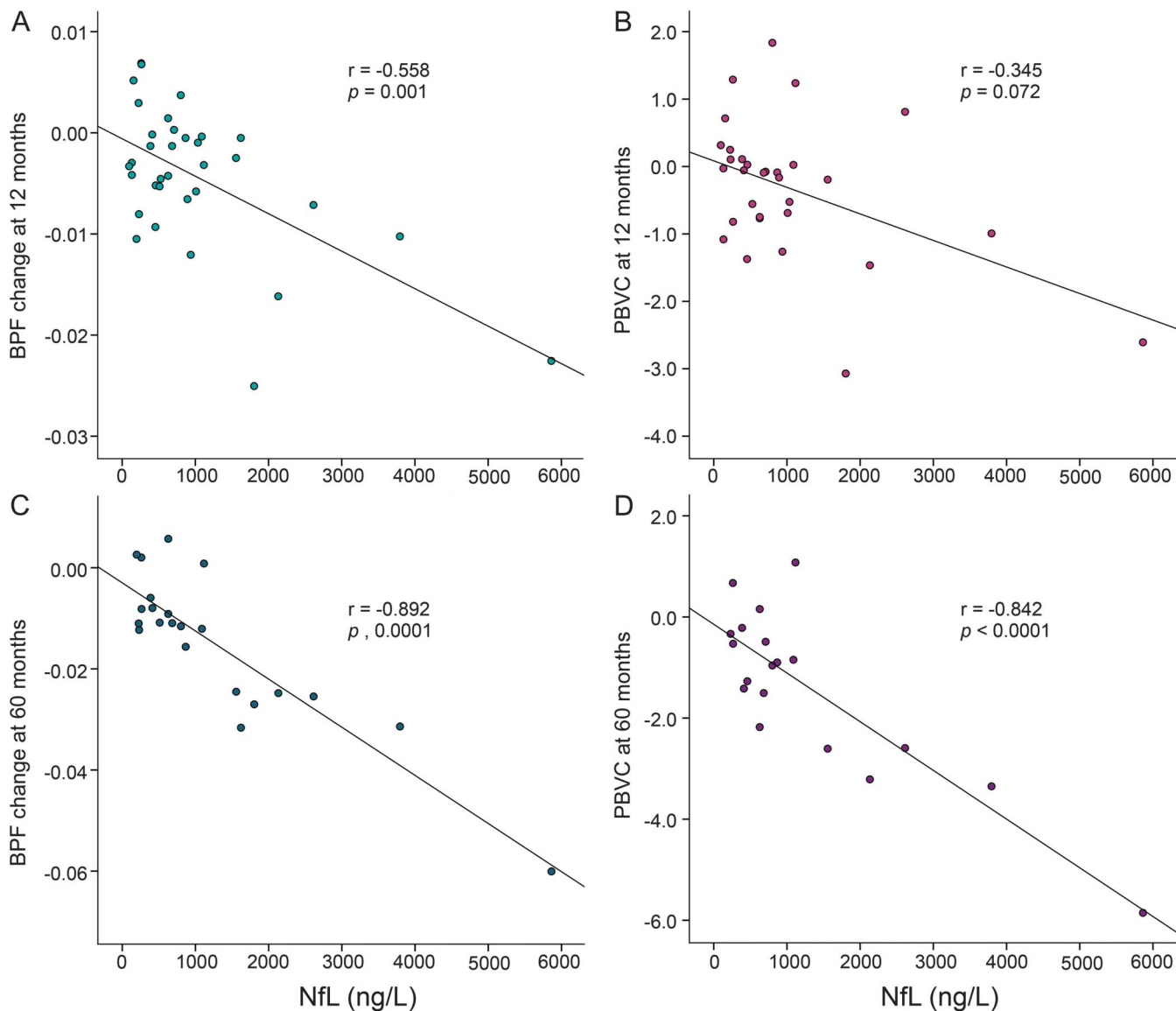
| | Baseline | | | Year 1 | | | | | | Year 5 | | | | | |
|------------------|----------|---------|---------|---------|---------|-------|-------|--------|--------|--------|-------|-------|-------|---------|---------|
| | T2LN | T2LV | Gd+ | T2LN | T2LV | T2New | Gd+ | BPFΔ | PBVC | T2LN | T2LV | T2New | Gd+ | BPFΔ | PBVC |
| No. ^a | 39 | 39 | 39 | 38 | 38 | 40 | 38 | 33 | 30 | 21 | 21 | 23 | 22 | 21 | 18 |
| r^b | 0.595 | 0.617 | 0.538 | 0.619 | 0.649 | 0.492 | 0.456 | -0.558 | -0.345 | 0.570 | 0.609 | 0.589 | 0.167 | -0.892 | -0.842 |
| p Value | <0.0001 | <0.0001 | <0.0001 | <0.0001 | <0.0001 | 0.001 | 0.004 | 0.001 | 0.072 | 0.007 | 0.003 | 0.003 | 0.456 | <0.0001 | <0.0001 |

Abbreviations: BPFΔ = brain parenchymal fraction percentage change; Gd+ = number of gadolinium-enhancing lesions; PBVC = percentage brain volume change; T2LN = T2 lesion number; T2LV = T2 lesion volume; T2New = number of new T2 lesions.

^aData assessed in MRI scans performed from 2001.

^bMRI inflammatory measures assessed with Spearman correlation. BPFΔ and PBVC values correspond to partial correlations adjusted for age and Gd+ lesions at baseline.

Figure 2 Scatterplots showing the correlations between NfL levels and brain volume changes



The graphs represent the raw data. The correlation coefficients and p values correspond to the partial correlations adjusted for age and baseline gadolinium-enhancing lesions. (A) BPF change at 1 year. (B) PBVC at 1 year. (C) BPF change at 5 years. (D) PBVC at 5 years. BPF = brain parenchymal fraction; NfL = neurofilament light chain; PBVC = percentage brain volume change.

We observed significant associations between NfL and MRI inflammatory parameters on baseline and follow-up MRIs, with similar findings during the replication phase. Regarding Gd-enhancing lesions, in the screening, we found correlations with NfL levels at baseline and 1 year, observed again in the replication phase, but not at 5 years. There are previous reports of correlations with T2 lesion number and Gd-enhancing lesions in CIS and/or MS groups,^{18,25} unconfirmed in other CIS-specific studies.^{20,34} The most striking findings concern MRI neurodegenerative parameters: our results suggest baseline NfL levels increase not only in association with lesion load, accrual, and activity, but also independently of these parameters.^{35,36} On the contrary, a previous study did

not find any correlations between NfL levels and brain volume change in CIS.²⁰ Although the authors did adjust for age and used the SIENA software, they assessed the follow-up only at 1 year and did not control by Gd-enhancing lesions. We additionally evaluated the association between baseline NfL levels and brain volume changes at 5 years. Our findings are partly supported by those of Burman et al.,¹⁵ in which NfL levels were elevated irrespective of Gd-enhancing lesions in progressive MS. Unfortunately, our appraisal is limited: our sample size was small and we were unable to repeat this evaluation in the replication phase. However, by estimating both BPF Δ and PBVC, the latter a robust method, we believe our results are reliable and should be assessed in future studies.

Table 3 Uni- and multivariate analyses for CDMS and 2010 McDonald MS with NfL as a continuous variable

| | p Value | HR/aHR ^a | 95% CI |
|--|---------|---------------------|---------------|
| Conversion to CDMS | | | |
| Univariate analysis | | | |
| DMT ^b | 0.701 | 1.126 | 0.615-2.062 |
| Hospital | 0.063 | 1.722 | 0.971-3.052 |
| NfL-100 ^c | <0.0001 | 1.009 | 1.005-1.014 |
| OCBs | <0.0001 | 6.074 | 2.580-14.300 |
| T2 lesions 1-3 | 0.047 | 8.537 | 1.026-71.004 |
| T2 lesions ≥4 | 0.002 | 21.709 | 2.982-158.026 |
| Multivariate analysis^d | | | |
| NfL-100 | 0.040 | 1.005 | 1.000-1.011 |
| OCBs | 0.048 | 2.597 | 1.009-6.683 |
| T2 lesions 1-3 | 0.071 | 7.225 | 0.843-61.920 |
| T2 lesions ≥4 | 0.022 | 11.469 | 1.432-91.868 |
| 2010 McDonald criteria | | | |
| Univariate analysis | | | |
| DMT | 0.292 | 0.736 | 0.416-1.302 |
| NfL-100 | <0.0001 | 1.009 | 1.005-1.013 |
| Hospital ^e | 0.001 | 2.232 | 1.404-3.546 |
| OCBs | <0.0001 | 8.427 | 4.189-16.951 |
| T2 lesions 1-3 | 0.020 | 11.842 | 1.480-94.741 |
| T2 lesions ≥4 | <0.0001 | 52.103 | 7.220-375.983 |
| Multivariate analysis^f | | | |
| NfL-100 | 0.155 | 1.004 | 0.999-1.008 |
| OCBs | 0.012 | 2.669 | 1.236-5.762 |
| T2 lesions 1-3 | 0.036 | 9.593 | 1.165-79.034 |
| T2 lesions ≥4 | 0.002 | 25.676 | 3.347-196.974 |

Abbreviations: aHR = adjusted hazard ratio; CDMS = clinically definite multiple sclerosis; CI = confidence interval; DMT = disease-modifying treatment; HR = hazard ratio; NfL = neurofilament light chain; OCB = oligoclonal bands.

^a Univariate analyses: HR; multivariate analyses: aHR.

^b DMT: before conversion to CDMS or fulfillment of the 2010 McDonald criteria.

^c HR increase for every 100 ng/L.

^d Results were not modified before and after adjusting for DMT.

^e Higher risk of McDonald MS for cases from Ramón y Cajal Hospital.

^f Results adjusted for hospital and DMT. When only adjusting for hospital, aHR for NfL remained similar (aHR 1.004, 95% CI 0.999-1.009) but with a *p* value of 0.081.

Of note, there was a significant correlation between NfL and NfH levels, in accordance with previous publications on CIS and relapsing-remitting MS,^{18,20,31} although contrary to our results, some of these studies demonstrated associations between NfH and EDSS or MRI parameters.^{18,20} Thus, the origin of this correlation in our study is a matter of debate and deserves further study.

In the replication phase, considering that a dichotomic biomarker could be more useful in the clinical practice, we explored a 900 ng/L cutoff,²⁵ observing that more NfL-positive than -negative patients evolved

to CDMS and McDonald MS. Two other studies explored cutoff values to either identify relapses or predict evolution to secondary progressive MS.^{13,37} Their cutoff values, however, are lower than ours, a finding possibly related not only to the different outcomes but also to the use of a different assay,¹³ or the calculation methods based on NfL levels in patients with MS,^{13,37} whereas we established our cutoff using a control group with noninflammatory neurologic diseases under the rationale that this value would be more stable across studies. To assess whether this is a better approach, the cutoff should be evaluated in other cohorts. One study explored NfL levels and fulfillment of the 2010 dissemination in space criteria in optic neuritis,³⁴ whereas we used the cutoff value to investigate the risk of developing both CDMS and 2010 McDonald MS at the time of a CIS. Another contribution is the assessment of NfL as an independent risk factor for MS. In the univariate analyses, we found a higher risk of CDMS and McDonald MS for every 100-ng/L increase in NfL levels, maintained for CDMS in the multivariate analysis. The nonsignificant results for McDonald MS could be conditioned by the distribution of patients over the different categories in the multivariate analysis, the correlations between NfL levels and T2 lesions, and the strong predictive value of T2 lesions for dissemination in space fulfillment. When assessing NfL status, the increased risk of both CDMS and McDonald MS in the univariate analyses was lost in the multivariate models. Thus, although NfL status could be more practical in the clinical setting, in our study, only NfL as a continuous variable is an independent risk factor for CDMS. Nevertheless, the value of NfL should be considered in context with OCB and T2 lesion results and, according to a recent study, it could be considered a weak risk factor for MS.⁶

Despite conducting measures to minimize bias, a limitation in the replication phase concerns the NfL determination in 2 different centers. Although the corresponding intra- and interplate CVs were within the assay's standardized limits, we did not perform quality controls such as testing the same samples in both centers.³⁸ Besides, the shorter follow-up of the Vall d'Hebron cohort in the replication phase may preclude excluding patients with potential alternative diagnoses; nevertheless, the specific inclusion of typical CIS presentations lowers this risk.⁶ Furthermore, a shortcoming of NfL levels is their determination in CSF because some centers do not perform lumbar punctures routinely. It is impossible to quantify serum NfL levels with the available ELISA^{29,30}; however, a new electrochemiluminescence array showed promising results in detecting higher serum NfL levels in patients with CIS compared to healthy controls.³⁹

Our biomarker list was probably not thoroughly comprehensive. We did, however, search for updates

throughout the study duration. Besides, the properties of the selected assays could have influenced our findings. Another limitation is the difficulty in assessing all biomarkers in the same patients without compromising the sample size given the limited CSF aliquot availability. Nevertheless, the homogeneous data collection and systematic sample storage procedures could aid to minimize bias. Finally, we did not compare the NfH electrochemiluminescence assay to other methods.

Therefore, elevated NfL levels indicate a higher risk of evolving to MS. However, the increase in risk for every 100 ng/L could be difficult to interpret in the daily clinical practice and, compared to OCBs and T2 lesions, NfL levels are a weak risk factor for MS. The search for a cutoff value should be assessed in other cohorts and, finally, the value of NfL as a marker of axonal damage may be more relevant as suggested by the correlations with medium-term brain volume loss.

AUTHOR CONTRIBUTIONS

Georgina Arrambide participated in the design of the work and on the acquisition, statistical analysis, and interpretation of data for the work, drafted and revised it, approved the present version to be published, and agreed to be accountable for all aspects of the work. Carmen Espejo participated in the design of the work and on the analysis and interpretation of data for the work, drafted and revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Herena Eixarch participated in the analysis and interpretation of data for the work, drafted and revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Luisa M. Villar participated in the acquisition, analysis, and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. José C. Alvarez-Cermeño participated in the acquisition, analysis, and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Carmen Picón participated in the acquisition and analysis of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Jens Kuhle participated in the acquisition, analysis, and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Giulio Disanto participated in the acquisition and analysis of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Ludwig Kappos participated in the analysis and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Jaume Sastre-Garriga participated in the design and acquisition, analysis, and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Deborah Pareto participated in the acquisition, analysis, and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Eva Simon participated in the acquisition of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Manuel Comabella participated in the acquisition and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Jordi Río participated in the acquisition and interpretation of data for the

work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Carlos Nos participated in the interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Carmen Tur participated in the acquisition and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Joaquín Castelló participated in the acquisition and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Angela Vidal-Jordana participated in the acquisition and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Ingrid Galán participated in the acquisition and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. María J. Arévalo participated in the acquisition of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Cristina Auger participated in the analysis and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Alex Rovira participated in the analysis and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Xavier Montalban participated in the conception and design of the work and on the acquisition, analysis, and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work. Mar Tintore participated in the conception and design of the work and on the acquisition, analysis, and interpretation of data for the work, revised it for important intellectual content, approved the present version to be published, and agreed to be accountable for all aspects of the work.

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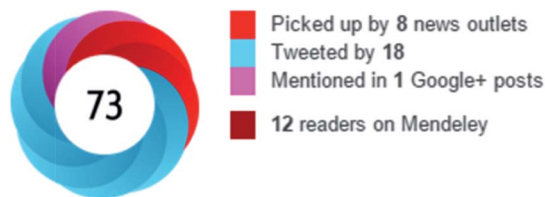
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Neurofilament light chain level is a weak risk factor for the development of MS

Georgina Arrambide, Carmen Espejo, Herena Eixarch, et al.

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generates the question: Would this high exacerbation rate happen if patients were allowed to achieve remission with steroids without using MMF?

It was previously illustrated that among patients with MG who were treated with steroids, 80.2% (93/116) achieved remission/minimal manifestations status and only 18% (17/93) experienced an exacerbation afterward.² This clues in that the high exacerbation rate reported by Oskarsson et al. after discontinuing MMF may not reflect the natural disease activity but rather reflects MMF dependence, and that continuing MMF mainly treats the dependence. Both studies are limited by a lack of information on the prednisone dose required to maintain remission/minimal manifestations status (an important data point to determine whether MMF is effective).

Author Response: Björn Oskarsson, Sacramento; David M. Rocke, Davis; Karsten Dengel, Sacramento; David P. Richman, Davis, CA: We thank Dimachkie et al. for their interest in our article.¹ However, in opposition to Dimachkie et al., we would consider MG exacerbations after discontinuation of MMF as MG exacerbations rather than a novel MMF dependence

condition. Given this premise, patients with pharmacologically controlled MG seem a more appropriate control group compared to patients who had MG but sustain remission without pharmacologic treatment. We would suspect that patients without symptoms or treatment may have a less-active disease compared to a population requiring treatment. This last group is also rare in our clinic, further making such a comparison less meaningful.

The patients were selected on the basis of being on stable doses of prednisone (0–25 mg/d; see table 2)¹ and only MMF was varied. Our article does not address corticosteroid treatment of MG and we do not argue that corticosteroids are not an effective treatment of MG. The interest in treating MG with MMF stems primarily from MMF's more favorable side-effect profile.

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CORRECTIONS

Neurofilament light chain level is a weak risk factor for the development of MS

In the article “Neurofilament light chain level is a weak risk factor for the development of MS” by G. Arrambide et al.,¹ there is an error in figure 2. In panel C, the p value should read “ $p < 0.0001$ ” rather than “ $p, 0.0001$ ” as originally published. The editorial staff regrets the error.

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1. Arrambide G, Espejo C, Eixarch H, et al. Neurofilament light chain level is a weak risk factor for the development of MS. *Neurology* 2016;87:1076–1084.

Pediatric multiple sclerosis: Conventional first-line treatment and general management

In the article “Pediatric multiple sclerosis: Conventional first-line treatment and general management” by A. Ghezzi et al.,¹ there is an error in the fourth author's name, which should have read “Teri Schreiner” rather than “Teri Shreiner” as originally published. The authors regret the error.

REFERENCE

1. Ghezzi A, Amato MP, Makhani N, et al. Pediatric multiple sclerosis: conventional first-line treatment and general management. *Neurology* 2016;87(suppl 2):S97–S102.

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